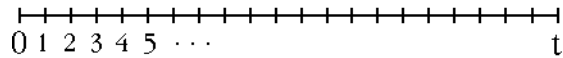


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Contents

- ☞ Poisson process as limiting case of the Bernoulli process
- ☞ Poisson distribution
- ☞ Exponential distribution
- ☞ Erlang (Gamma) distribution

Consider the following situation:



A time interval of length t seconds is divided into one-second intervals, with the probability of a vehicle arriving at an intersection during a one-second interval being a small number p . (Assume that the probability that more than one vehicle arrives is negligible.)

Consider the Bernoulli process $\{X_k; k=1,2,\dots\}$ where $X_k = 1$ if a vehicle arrives during the k^{th} second, and the associated counting process $\{N_t\}$ which counts the number of arrivals during the interval $[0,t]$.

Then N_t has the binomial distribution:

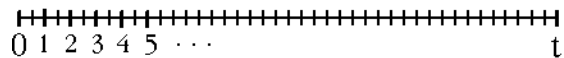
$$P(N_t = x) = \binom{t}{x} p^x (1-p)^{t-x}$$

with expected value $v = tp$.

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$$\begin{aligned} P(N_t = x) &= \binom{t}{x} \left(\frac{v}{t}\right)^x \left(1 - \frac{v}{t}\right)^{t-x} \\ &= \frac{t!}{x!(t-x)!} \left(\frac{v}{t}\right)^x \left(1 - \frac{v}{t}\right)^{t-x} \\ &= \frac{v^x}{x!} \left(1 - \frac{v}{t}\right)^t \frac{t!}{(t-x)!} \frac{1}{t^x \left(1 - \frac{v}{t}\right)^x} \end{aligned}$$

Consider what happens as we divide $[0,t]$ into n smaller time intervals, but in such a way that the expected number of arrivals in $[0,t]$ remains constant, v .



That is, the probability of an arrival in each of these small intervals must be $\frac{v}{n}$, and

$$P(N_t = x) = \binom{n}{x} \left(\frac{v}{n}\right)^x \left(1 - \frac{v}{n}\right)^{n-x}$$

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Consider the limit of this distribution as $n \rightarrow +\infty$

$$\begin{aligned} P(N_t = x) &= \frac{v^x}{x!} \left(1 - \frac{v}{n}\right)^n \frac{n!}{(n-x)!} \frac{1}{n^x \left(1 - \frac{v}{n}\right)^x} \\ &\quad \downarrow \quad \parallel \\ &= e^{-v} \frac{n(n-1)(n-2) \dots (n-x+1)}{\left[n \left(1 - \frac{v}{n}\right)\right]^x} \\ &\quad \rightarrow \frac{n^x}{n^x} = 1 \end{aligned}$$

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$$P(N_t = x) = \frac{v^x}{x!} e^{-v}$$

If the arrival rate is λ /second, then $v = \lambda t$ and

$$P(N_t = x) = \frac{(\lambda t)^x}{x!} e^{-\lambda t}$$

for $x=0, 1, 2, 3, \dots$

Poisson Distribution

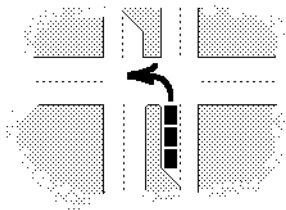


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Example

A left-turn lane at an intersection has a capacity of 3 autos. 30% of autos arriving at the intersection wish to turn left. The *expected number* of autos arriving during a red signal is 6.

What is the probability that the capacity of the left-turn lane is exceeded during a red signal?



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$$P\{X > 3 | N \text{ arrivals}\} = \sum_{x=4}^N \binom{N}{x} (0.3)^x (0.7)^{N-x}$$

binomial distr.

$$= 1 - \sum_{x=0}^3 \binom{N}{x} (0.3)^x (0.7)^{N-x}$$

$$P\{N \text{ arrivals}\} = \frac{6^N}{N!} e^{-6}$$

Poisson distr.

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N	P(N)	P(X N)	P(X N)P(N)
0	0.00247875	0.00000000	0.00000000
1	0.01487251	0.00000000	0.00000000
2	0.04461754	0.00000000	0.00000000
3	0.08923508	0.00000000	0.00000000
4	0.13385262	0.00810000	0.00108421
5	0.16062314	0.03078000	0.00494398
6	0.16062314	0.07047000	0.01131941
7	0.13767698	0.12603600	0.01735226
8	0.10325773	0.19410435	0.02004278
9	0.06883849	0.27034090	0.01860986
10	0.04130309	0.35038928	0.01447216
11	0.02252896	0.43043756	0.00969731
12	0.01126448	0.50748423	0.00571655
13	0.00519899	0.57939435	0.00301227
14	0.00222814	0.64483257	0.00143678
15	0.00089126	0.70313207	0.00062667
⋮	⋮	⋮	⋮
			0.1083

The probability that the capacity of the left-turn lane is exceeded during each red signal is about 11%



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Poisson Distribution

$$P(N_t = x) = \frac{(\lambda t)^x}{x!} e^{-\lambda t}$$

for $x=0, 1, 2, 3, \dots$

Mean Value

$$E(N_t) = \lambda t$$

Variance

$$\text{Var}(N_t) = \lambda t$$

mean and variance are equal!



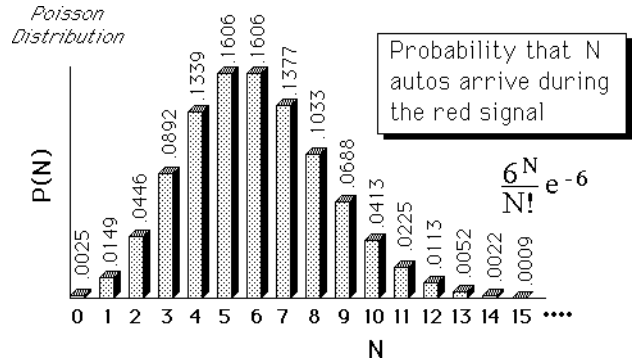
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Given that N autos arrive, the number X of left-turning autos has the *binomial* distribution.

The number N of autos arriving during the red signal has the *Poisson* distribution.

$$P\{X > 3\} = \sum_{N=4}^{\infty} \underbrace{P\{X > 3 | N \text{ arrivals}\}}_{\text{computed using binomial distr.}} \underbrace{P\{N \text{ arrivals}\}}_{\text{computed using Poisson distr.}}$$

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Time between arrivals

Suppose that the number of arrivals in an interval has the Poisson distribution with arrival rate λ /second.

Let T_1 = time of the first arrival. What is the distribution of T_1 ?



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$$P\{T_1 > t\} = P\{\text{NO arrivals occur in interval } [0, t]\}$$

$$= \frac{(\lambda t)^0}{0!} e^{-\lambda t} = e^{-\lambda t}$$

← Poisson distribution $P\{N_t = 0\}$

$$\text{CDF: } P\{T_1 \leq t\} = F(t) = 1 - e^{-\lambda t}$$

$$\text{Density function: } f(t) = \frac{d}{dt} F(t) = \lambda e^{-\lambda t}$$

Exponential Distribution

Exponential Distribution

$$F(t) = 1 - e^{-\lambda t}$$

Mean Value

$$E(T_1) = \frac{1}{\lambda}$$

Variance

$$\text{Var}(T_1) = \frac{1}{\lambda^2}$$

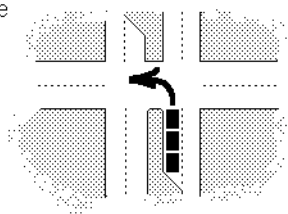
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Example

Suppose that the arrival rate for northbound autos is 6 per 30 second red signal, i.e., 0.2/second

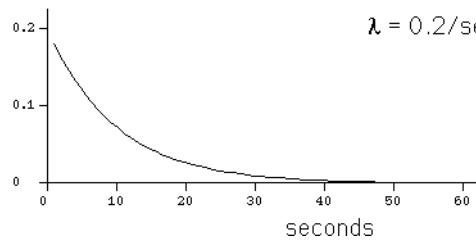
What is the distribution of the arrival time of the first auto?
(This will also be the distribution of the time between arrivals!)



$$f(t) = \lambda e^{-\lambda t}$$

Exponential Distribution

$$\lambda = 0.2/\text{sec.}$$



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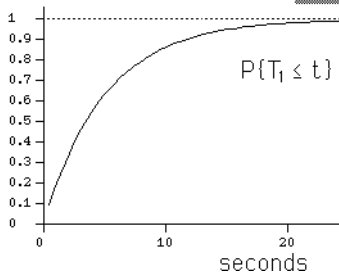
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$$\lambda = 0.2/\text{sec.}$$

t	F(t)
1	0.18127
2	0.32968
3	0.45119
4	0.55067
5	0.63212
6	0.69881
7	0.75340
8	0.79810
9	0.83470
10	0.86466
11	0.88920
12	0.90928
13	0.92573
14	0.93919
15	0.95021
16	0.95924
17	0.96663
18	0.97268
19	0.97763
20	0.98168

$$F(t) = 1 - e^{-\lambda t}$$

Exponential Distribution



Memoryless Property

Exponential Distribution

Suppose that it is known that, at time t_0 , the first arrival has not yet occurred, i.e., $T_1 > t_0$.

What is the conditional distribution of T_1 ?

That is, what is $P\{T_1 \leq t \mid T_1 > t_0\}$ for $t > t_0$?



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Memoryless Property

Exponential Distribution

$$\begin{aligned} P\{T_1 \leq t \mid T_1 > t_0\} &= \frac{P\{T_1 \leq t \cap T_1 > t_0\}}{P\{T_1 > t_0\}} = \frac{P\{t_0 \leq T_1 \leq t\}}{P\{T_1 > t_0\}} \\ &= \frac{F(t) - F(t_0)}{1 - F(t_0)} = \frac{(1 - e^{-\lambda t}) - (1 - e^{-\lambda t_0})}{e^{-\lambda t_0}} \\ &= \frac{e^{-\lambda t_0} - e^{-\lambda t}}{e^{-\lambda t_0}} = 1 - e^{-\lambda(t-t_0)} \end{aligned}$$

Memoryless Property

Exponential Distribution

$$P\{T_1 \leq t \mid T_1 > t_0\} = 1 - e^{-\lambda(t-t_0)} = P\{T_1 \leq t - t_0\}$$

If the time τ is reckoned from time t_0 , i.e., $\tau = t - t_0$, then

$$P\{T_1 \leq t \mid T_1 > t_0\} = P\{T_1 \leq t - t_0\} = P\{T_1 \leq \tau\}$$

In other words, the failure of an arrival to occur before time t_0 does not alter one's prediction of the length of time (from t_0) before the next arrival.

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Time of kth Arrival

Let T_k = time of kth arrival,
 $\tau_k = T_k - T_{k-1}$ = time between arrivals k-1 and k.
 Suppose that τ_k (k=1,2,3,...) have identical and independent exponential distributions with rate λ .
 Then T_k is the *sum* of k random variables with exponential distributions.
 It is said to have a **k-Erlang** distribution.



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Erlang Distribution

CDF $F(t) = \frac{\Gamma(k, \lambda t)}{\Gamma(k)}$

where $\Gamma(k, x)$ is the *"incomplete Gamma function"* defined by

$$\Gamma(k, x) = \int_0^x e^{-u} u^{k-1} du$$

tabulated

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Erlang Distribution

Mean Value $\mu = \frac{k}{\lambda}$ *(These expressions result from the fact that the random variable is the sum of k i.i.d. random variables.)*

Variance $\sigma^2 = \frac{k}{\lambda^2}$

More generally, when k is not an integer, the probability distribution is called the **Gamma** distribution.

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Bernoulli process

Poisson process

Binomial distn. # of events

Poisson distn.

Geometric distn. time until 1st event

Exponential distn.

Pascal distn. time until kth event

Erlang distn.



Erlang Distribution

time of kth arrival in a Poisson process

Density function
 (k>0, λ>0, t≥0)

$$f(t) = \frac{\lambda (\lambda t)^{k-1} e^{-\lambda t}}{(k-1)!}$$

$$= \frac{\lambda (\lambda t)^{k-1} e^{-\lambda t}}{\Gamma(k)}$$

where the **Gamma** function is defined by
 (for k>0, not necessarily integer!)

$$\Gamma(k) = \int_0^\infty e^{-u} u^{k-1} du$$

$$= (k-1)! \quad \text{if } k \text{ integer}$$

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Alternate computation, when k is integer

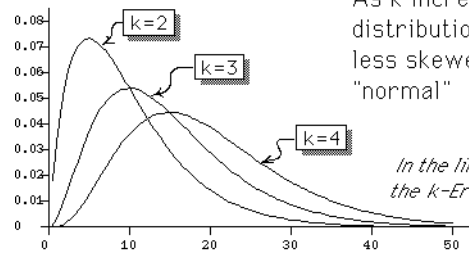
CDF $F(t) = P\{T_k \leq t\} = P\{N_t \geq k\}$
 $= 1 - P\{N_t < k\}$
 $= 1 - P\{N_t \leq k-1\}$
 where N_t = # arrivals at time t has the **Poisson** distribution:

$$F(t) = 1 - \sum_{x=0}^{k-1} \frac{(\lambda t)^x}{x!} e^{-\lambda t}$$

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Erlang Distribution

example:
 $\lambda = 0.2$



As k increases, the distribution becomes less skewed, and more "normal"

In the limit, as k → ∞, the k-Erlang distribution converges to the Normal distribution!

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