

# Extreme Value Distributions

## The Weibull Distribution

© D. L. Bricker  
Dept of Mechanical & Industrial Engineering  
The University of Iowa

# Asymptotic Distributions

From the *Central Limit Theorem* we know that, for “large”  $n$ ,

$Y = \sum_{i=1}^n X_i$  has approximately a *Normal* distribution

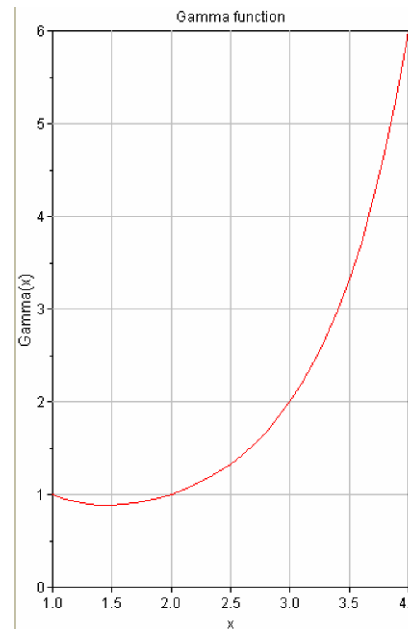
$Y = \prod_{i=1}^n X_i$  has approximately a *Lognormal* distribution

$$f_Y(y) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left\{-\frac{1}{2}\left(\frac{y-\mu}{\sigma}\right)^2\right\} \quad f_Y(y) = \frac{1}{y\sigma_x\sqrt{2\pi}} \exp\left\{-\frac{1}{2}\left(\frac{\ln(y/\mu)}{\sigma}\right)^2\right\}$$

When the *minimum* value of  $X$  is 0, i.e.,  $X$  is nonnegative, the limiting distribution for  $T = \min\{X_i\}$  as  $n \rightarrow \infty$  is the **Weibull** distribution:

## Asymptotic Distributions

|                           |  |
|---------------------------|--|
| <b>CDF</b>                | $F_T(t) = 1 - e^{-(t/u)^k}$  |
| <b>Mean value</b>         | $\mu_T = u\Gamma\left(1 + \frac{1}{k}\right)$  |
| <b>Standard deviation</b> | $\sigma_T = u\sqrt{\Gamma\left(1 + \frac{2}{k}\right) - \Gamma^2\left(1 + \frac{1}{k}\right)}$ |



The “Gamma” function  $\Gamma$  is a generalization of the *factorial* function, defined for *all*  $x \geq 0$ , (not necessarily integer) and satisfies  $\Gamma(1+x) = x!$  when  $x$  is a nonnegative integer.

Table of values of  $\Gamma\left(1+\frac{1}{k}\right)$  for  $k = 0.1$  through 9.9

|   | 0.0      | 0.1      | 0.2     | 0.3    | 0.4    | 0.5    | 0.6    | 0.7    | 0.8    | 0.9    |
|---|----------|----------|---------|--------|--------|--------|--------|--------|--------|--------|
| 0 | $\infty$ | 3.63E+06 | 120.000 | 9.2610 | 3.3230 | 2.0000 | 1.5050 | 1.2660 | 1.1330 | 1.0520 |
| 1 | 1.0000   | 0.9649   | 0.9407  | 0.9236 | 0.9114 | 0.9027 | 0.8966 | 0.8922 | 0.8893 | 0.8874 |
| 2 | 0.8862   | 0.8857   | 0.8856  | 0.8859 | 0.8865 | 0.8873 | 0.8882 | 0.8893 | 0.8905 | 0.8917 |
| 3 | 0.8930   | 0.8943   | 0.8957  | 0.8970 | 0.8984 | 0.8997 | 0.9011 | 0.9025 | 0.9038 | 0.9051 |
| 4 | 0.9064   | 0.9077   | 0.9089  | 0.9102 | 0.9114 | 0.9126 | 0.9137 | 0.9149 | 0.9160 | 0.9171 |
| 5 | 0.9182   | 0.9192   | 0.9202  | 0.9213 | 0.9222 | 0.9232 | 0.9241 | 0.9251 | 0.9260 | 0.9269 |
| 6 | 0.9277   | 0.9286   | 0.9294  | 0.9302 | 0.9310 | 0.9318 | 0.9325 | 0.9333 | 0.9340 | 0.9347 |
| 7 | 0.9354   | 0.9361   | 0.9368  | 0.9375 | 0.9381 | 0.9387 | 0.9394 | 0.9400 | 0.9406 | 0.9412 |
| 8 | 0.9417   | 0.9423   | 0.9429  | 0.9434 | 0.9439 | 0.9445 | 0.9450 | 0.9455 | 0.9460 | 0.9465 |
| 9 | 0.9470   | 0.9474   | 0.9479  | 0.9484 | 0.9488 | 0.9493 | 0.9497 | 0.9501 | 0.9505 | 0.9509 |

For example,  $\Gamma\left(1+\frac{1}{2.5}\right) = \Gamma(1.4) = 0.8873$

Table of values of  $\Gamma\left(1+\frac{2}{k}\right)$  for  $k = 0.1$  through 9.9

|   | 0.0    | 0.1      | 0.2     | 0.3    | 0.4    | 0.5     | 0.6    | 0.7    | 0.8    | 0.9    |
|---|--------|----------|---------|--------|--------|---------|--------|--------|--------|--------|
| 0 | 0      | 2.43E+18 | 3628800 | 2594   | 120    | 24.0000 | 9.2610 | 5.0290 | 3.3230 | 2.4790 |
| 1 | 2.0000 | 1.7020   | 1.5050  | 1.3660 | 1.2660 | 1.1910  | 1.1330 | 1.0880 | 1.0520 | 1.0230 |
| 2 | 1.0000 | 0.9808   | 0.9649  | 0.9517 | 0.9407 | 0.9314  | 0.9236 | 0.9170 | 0.9114 | 0.9067 |
| 3 | 0.9027 | 0.8994   | 0.8966  | 0.8942 | 0.8922 | 0.8906  | 0.8893 | 0.8882 | 0.8874 | 0.8867 |
| 4 | 0.8862 | 0.8859   | 0.8857  | 0.8856 | 0.8856 | 0.8857  | 0.8859 | 0.8862 | 0.8865 | 0.8868 |
| 5 | 0.8873 | 0.8877   | 0.8882  | 0.8887 | 0.8893 | 0.8899  | 0.8905 | 0.8911 | 0.8917 | 0.8923 |
| 6 | 0.8930 | 0.8936   | 0.8943  | 0.8950 | 0.8957 | 0.8963  | 0.8970 | 0.8977 | 0.8984 | 0.8991 |
| 7 | 0.8997 | 0.9004   | 0.9011  | 0.9018 | 0.9025 | 0.9031  | 0.9038 | 0.9044 | 0.9051 | 0.9058 |
| 8 | 0.9064 | 0.9070   | 0.9077  | 0.9083 | 0.9089 | 0.9096  | 0.9102 | 0.9108 | 0.9114 | 0.9120 |
| 9 | 0.9126 | 0.9132   | 0.9137  | 0.9143 | 0.9149 | 0.9154  | 0.9160 | 0.9166 | 0.9171 | 0.9176 |

For example,  $\Gamma\left(1+\frac{2}{2.5}\right) = \Gamma(1.8) = 0.9314$

## Computing Weibull Parameters

Given  $\mu_T$  and  $\sigma_T$ , we wish to solve for the parameters  $u$  &  $k$ :

$$\begin{cases} \mu_T = u\Gamma\left(1+\frac{1}{k}\right) \\ \sigma_T = u\sqrt{\Gamma\left(1+\frac{2}{k}\right) - \Gamma^2\left(1+\frac{1}{k}\right)} \end{cases}$$

This is a system of two nonlinear equations in two unknowns  $u$  &  $k$ , which might be solved by (for example) the *Newton-Raphson* method.

Solve for  $u$  in terms of the mean  $\mu_T$ :

$$\mu_T = u\Gamma\left(1+\frac{1}{k}\right) \Rightarrow u = \frac{\mu_T}{\Gamma\left(1+\frac{1}{k}\right)}$$

Use this to eliminate  $u$  from the second equation:

$$\sigma_T = u\sqrt{\Gamma\left(1+\frac{2}{k}\right) - \Gamma^2\left(1+\frac{1}{k}\right)}$$

$$\Rightarrow \sigma_T = \frac{\mu_T}{\Gamma\left(1+\frac{1}{k}\right)}\sqrt{\Gamma\left(1+\frac{2}{k}\right) - \Gamma^2\left(1+\frac{1}{k}\right)}$$

This gives us a *single* (nonlinear) equation in a *single* variable  $k$ .

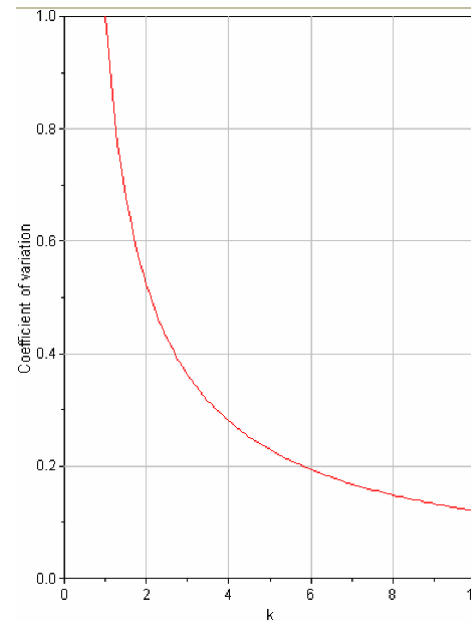
$$\sigma_T = \frac{\mu_T}{\Gamma\left(1+\frac{1}{k}\right)} \sqrt{\Gamma\left(1+\frac{2}{k}\right) - \Gamma^2\left(1+\frac{1}{k}\right)}$$

$$\Rightarrow \frac{\sigma_T}{\mu_T} = \frac{1}{\Gamma\left(1+\frac{1}{k}\right)} \sqrt{\Gamma\left(1+\frac{2}{k}\right) - \Gamma^2\left(1+\frac{1}{k}\right)}$$

$$\Rightarrow \frac{\sigma_T}{\mu_T} = \sqrt{\frac{\Gamma\left(1+\frac{2}{k}\right)}{\Gamma^2\left(1+\frac{1}{k}\right)} - 1}$$

**Coefficient  
Of Variation**

Thus, the coefficient of variation of the Weibull distribution is determined by  $k$  alone.



**Coefficient  
of Variation**

*as a function of  
parameter  $k$*

**Coefficient of Variation,  
as a function of parameter  $k$**

|   | 0.0    | 0.1    | 0.2    | 0.3    | 0.4    | 0.5    | 0.6    | 0.7    | 0.8    | 0.9    |
|---|--------|--------|--------|--------|--------|--------|--------|--------|--------|--------|
| 0 | -----  | 429.8  | 15.843 | 5.4077 | 3.1409 | 2.2361 | 1.7581 | 1.4624 | 1.2605 | 1.1130 |
| 1 | 1.0000 | 0.9102 | 0.8369 | 0.7757 | 0.7238 | 0.6790 | 0.6399 | 0.6055 | 0.5749 | 0.5475 |
| 2 | 0.5227 | 0.5003 | 0.4798 | 0.4611 | 0.4438 | 0.4279 | 0.4131 | 0.3994 | 0.3866 | 0.3747 |
| 3 | 0.3635 | 0.3529 | 0.3430 | 0.3337 | 0.3248 | 0.3165 | 0.3085 | 0.3010 | 0.2939 | 0.2870 |
| 4 | 0.2805 | 0.2744 | 0.2684 | 0.2628 | 0.2573 | 0.2521 | 0.2471 | 0.2424 | 0.2378 | 0.2333 |
| 5 | 0.2291 | 0.2250 | 0.2210 | 0.2172 | 0.2135 | 0.2099 | 0.2065 | 0.2031 | 0.1999 | 0.1968 |
| 6 | 0.1938 | 0.1908 | 0.1880 | 0.1852 | 0.1826 | 0.1800 | 0.1774 | 0.1750 | 0.1726 | 0.1703 |
| 7 | 0.1680 | 0.1658 | 0.1637 | 0.1616 | 0.1596 | 0.1576 | 0.1557 | 0.1538 | 0.1519 | 0.1501 |
| 8 | 0.1484 | 0.1467 | 0.1450 | 0.1434 | 0.1418 | 0.1402 | 0.1387 | 0.1372 | 0.1357 | 0.1343 |
| 9 | 0.1329 | 0.1315 | 0.1302 | 0.1288 | 0.1275 | 0.1263 | 0.1250 | 0.1238 | 0.1226 | 0.1215 |

Several methods might be used to estimate the parameters  $u$  &  $k$  of the Weibull distribution

- (a) method of moments, i.e., matching the mean and standard deviation
- (b) linear regression (after transforming to linear form)
- (c) maximum likelihood method

Method (a) requires that we have sufficient data to compute the mean & standard deviation.

Methods (b) & (c) require a sample of observations of  $T$ .

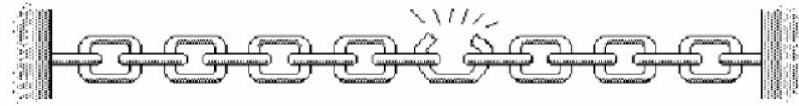
# Method of Moments

Given the coefficient of variation ( $\frac{\sigma}{\mu}$ ), we can either

- approximate  $k$  through the use of the table (or graph), or
- use a numerical method, e.g., the Newton-Raphson or Secant method, to solve the *nonlinear* equation

$$\frac{\sigma_T}{\mu_T} = \sqrt{\frac{\Gamma\left(1 + \frac{2}{k}\right)}{\Gamma^2\left(1 + \frac{1}{k}\right)} - 1}$$

*(The Newton-Raphson method requires derivatives, while the secant method does not!)*



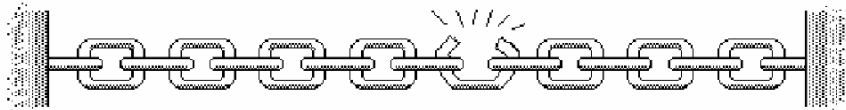
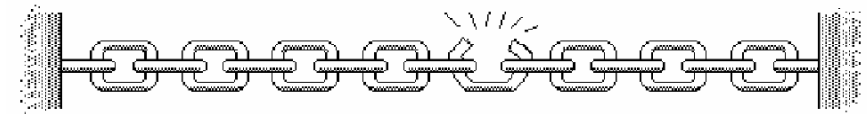
Let  $X_i$  = lifetime of link  $i$  of a chain ( $X_i \geq 0$ )

$T = \min\{X_1, X_2, \dots, X_n\}$  = lifetime of the chain

For large  $n$  (long chains) the distribution of  $T$  should be approximately Weibull.

Suppose that we have estimates for the mean  $\mu_T=150$  hours and standard deviation  $\sigma_T=50$  hours.

**E  
X  
A  
M  
P  
L  
E**



What is the probability that...

- the chain fails during its first 100 hours of use?
- the chain has not yet failed after 200 hours?

*Since the lifetime of the chain is the minimum of the lifetimes (times until failure) of the individual links of the chain, and these lifetimes are each bounded below by zero, we will assume the Weibull distribution.*

# Computation of the Parameter $k$

The *coefficient of variation* of the lifetime  $T$  is  $\frac{\sigma}{\mu} = \frac{50}{150} = \frac{1}{3}$ .

We will use the **Secant Method**

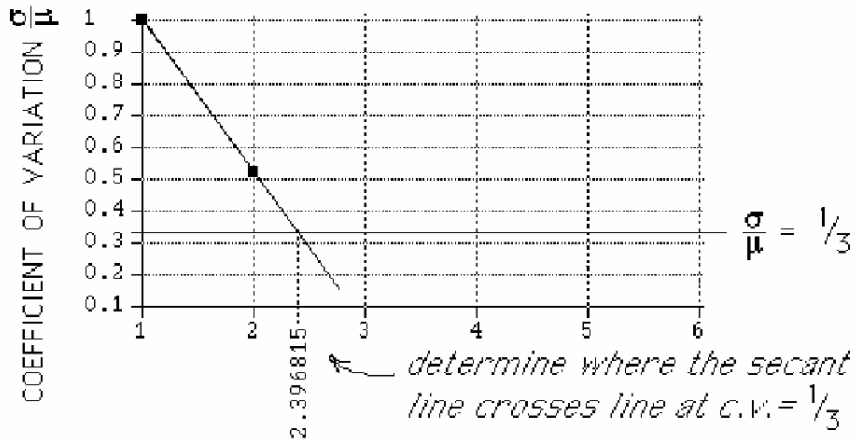
(which, unlike the Newton-Raphson method, does *not* require derivatives)

to solve the nonlinear equation  $\frac{\sigma_T}{\mu_T} = \sqrt{\frac{\Gamma\left(1 + \frac{2}{k}\right)}{\Gamma^2\left(1 + \frac{1}{k}\right)} - 1}$

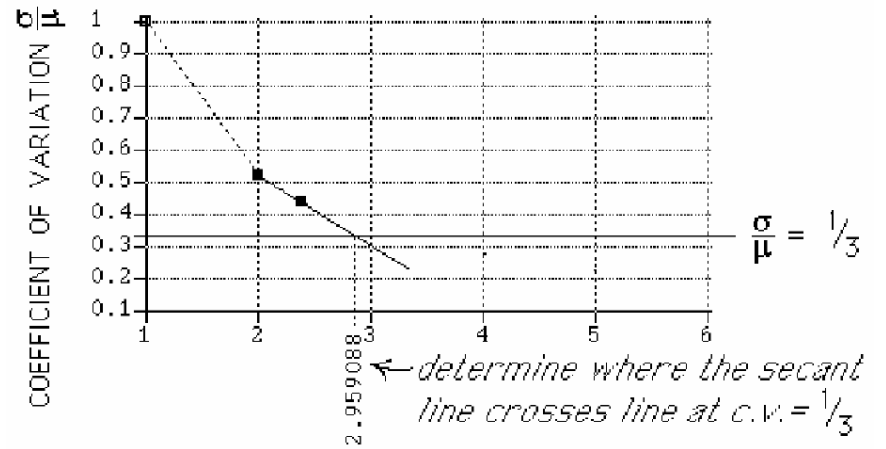
for  $k$ .

**Secant method**

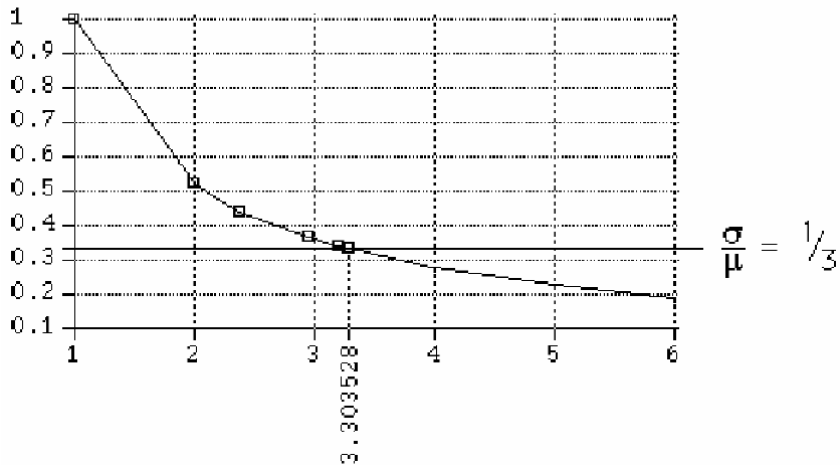
Start with "guesses"  $k=1$  &  $2$



Using last 2 values of  $k$ , draw a new secant line



Continue, until the procedure converges to  $k = 3.303528$



Using the secant method, we get the following approximations to the value of  $k$ :

| $k$      | error    |
|----------|----------|
| 0.250000 | 7.973291 |
| 2.000000 | 0.189390 |
| 2.042579 | 0.179577 |
| 2.821777 | 0.050615 |
| 3.127600 | 0.016787 |
| 3.279356 | 0.002206 |
| 3.302318 | 0.000109 |
| 3.303517 | 0.000001 |
| 3.303525 | 0.000000 |

Given  $k=3.3035$  and mean  $\mu_T = 150$ , we can now solve for the parameter  $u$ :

$$u = \frac{\mu_T}{\Gamma\left(1 + \frac{1}{k}\right)} = \frac{150}{\Gamma(1.3027)} = \frac{150}{0.8971} = 167.21$$

What is the probability that chain fails during its first 100 hours of use?  
i.e.,  $P\{T \leq 100\} = F_T(100) = ?$

What is the probability that it has not yet failed after 200 hours of use?  
i.e.,  $P\{T \geq 200\} = 1 - F_T(200) = ?$

$$F_T(t) = 1 - e^{-\left(\frac{t}{u}\right)^k} = 1 - e^{-\left(\frac{t}{167.21}\right)^{3.3025}}$$

Therefore,

$$\begin{aligned} P\{T \leq 100\} &= F_T(100) = 1 - \exp\left(\frac{100}{167.21}\right)^{3.3025} \\ &= 1 - \exp(0.60529)^{3.3025} = 1 - \exp(0.19052) = 1 - 0.82653 = 0.17347 \end{aligned}$$

That is, the chain is about 17% likely to fail during its first 100 hours of use.

$$F_T(t) = 1 - e^{-\left(\frac{t}{u}\right)^k} = 1 - e^{-\left(\frac{t}{167.21}\right)^{3.3025}}$$

Therefore,

$$P\{T \geq 200\} = 1 - F_T(200) = e^{-\left(\frac{200}{167.21}\right)^{3.3025}} = 0.16423$$

That is, the chain is about equally likely (16%) to survive its first 200 hours of use.