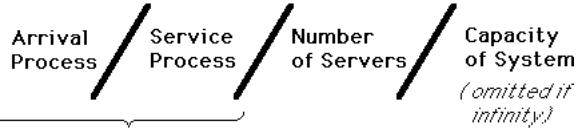
**Kendall's Notation**

M: Memoryless (Markovian)

E_k: Erlang-k

D: Deterministic

GI: General Interarrival times (but i.i.d.)

G: General service times (but i.i.d.)

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The "Memoryless" arrival process indicates a Poisson arrival process, in which the interarrival times have an *exponential* distribution.

Likewise, the "Memoryless" service process indicates that the service times have an *exponential* distribution.

Some Markovian Queues

- M/M/1
- M/M/c (c>1)
- M/M/1/N
- M/M/1/N/N



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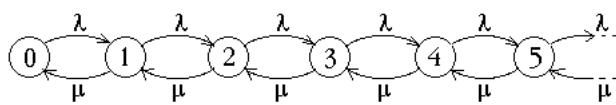
M/M/1 Interarrival times and service times both have exponential distributions,

with parameters λ & μ , respectively.
That is, the "customers" arrive at the rate of λ per unit time, and are served at the rate μ per unit of time.

It is assumed that the queue has infinite capacity, and that $\mu > \lambda$ (so that the queue length does not tend to increase indefinitely.)

In this case, it is possible to derive the probability distribution of the number of customers in the queueing system.

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M/M/1**Birth/Death Model**

$$\begin{aligned} \frac{1}{\pi_0} = 1 + \frac{\lambda}{\mu} + \left(\frac{\lambda}{\mu}\right)^2 + \left(\frac{\lambda}{\mu}\right)^3 + \left(\frac{\lambda}{\mu}\right)^4 + \dots \\ = \frac{1}{1 - \frac{\lambda}{\mu}} \quad (\text{Geometric Series}) \end{aligned}$$

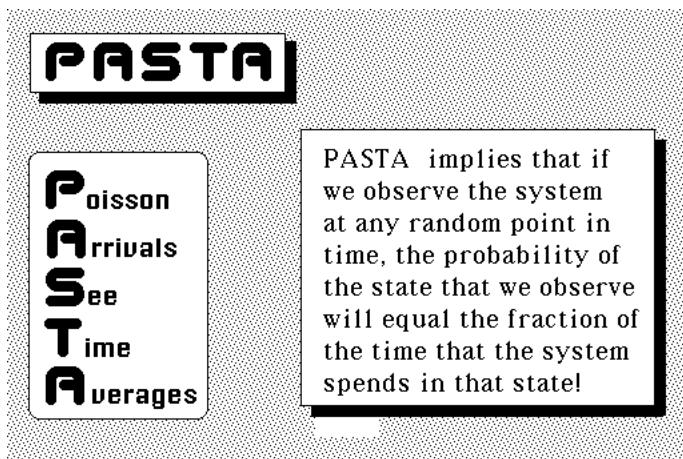
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M/M/1

$\pi = (\pi_0, \pi_1, \pi_2, \dots)$ denotes the "steady-state" distribution of the number of customers in this M/M/1 queueing system, i.e., 1+number in queue. Equivalently, π_i is the probability (in steady state) that an arriving customer will find i customers already in the queueing system.

$$\pi_i = \left(1 - \frac{\lambda}{\mu}\right) \left(\frac{\lambda}{\mu}\right)^i$$

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**M/M/1**

$$\pi_i = \left(1 - \frac{\lambda}{\mu}\right) \left(\frac{\lambda}{\mu}\right)^i$$

$$\pi_0 = 1 - \frac{\lambda}{\mu}$$

Probability that no customers are in the system (server is idle)

Probability that the server is busy

$$1 - \pi_0 = \frac{\lambda}{\mu} \equiv \rho < 1$$

utilization

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M/M/1

Average number of "customers" in the system (including any being served)

$$L = \sum_{j=0}^{\infty} j \pi_j = \sum_{j=0}^{\infty} j \underbrace{(1-\rho)}_{\pi_j} \rho^j = (1-\rho) \rho \sum_{j=0}^{\infty} j \rho^{j-1}$$

Next we will derive an expression for this infinite sum!

$$\text{where } \rho = \frac{\lambda}{\mu} < 1$$

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Derivative of a Geometric Series

This is not a geometric series, because of the appearance of the factor "j", but it is the sum of the derivatives of the terms of a geometric series!

$$\sum_{j=0}^{\infty} j \rho^{j-1} = \frac{d}{d\rho} \sum_{j=0}^{\infty} \rho^j = \frac{d}{d\rho} \left(\frac{1}{1-\rho} \right) = \frac{1}{(1-\rho)^2}$$

Geometric series,
with sum $\left(\frac{1}{1-\rho}\right)$

sum of derivatives equals derivative of the sum!

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M/M/1

Average number of "customers" in the system (including any being served)

$$L = \frac{\frac{\lambda}{\mu}}{1 - \frac{\lambda}{\mu}} = \frac{\rho}{1 - \rho}$$

This formula ONLY for M/M/1 queue!

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LITTLE's Queueing Formula

$$L = \lambda W$$

average number in the queueing system average arrival rate average time in system per customer

↳ applies to **any** queueing system having a steady state distribution (**whether Markovian or not!**)

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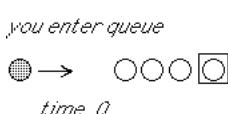
LITTLE's Queueing Formula

$$L = \lambda W$$

Intuitive argument:

Suppose that you join a queue and spend W minutes before you have been served and leave.

During those W minutes, customers have been arriving and joining the queue behind you at the average rate of λ per minute. Thus, when you are ready to leave, you should expect to see λW customers remaining in the system behind you.



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M/M/1

For the M/M/1 queueing system, Little's formula implies that

$$W = \frac{L}{\lambda} = \frac{\rho}{\lambda(1-\rho)}$$

where $\rho = \frac{\lambda}{\mu}$

$$\Rightarrow W = \frac{1}{\mu - \lambda}$$

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Denote

L_q = average number of customers in the queue (not including those being served)

W_q = average time spent per customer in the queue (not including service time)

Then for any queueing system having a steady state distribution,

$$W_q = W - \frac{1}{\mu} \quad \text{average rate of service}$$

$$L_q = \lambda W_q \quad \text{Little's Formula}$$

Thus, given any one of the quantities

L , W , L_q , and W_q

we can compute the others.

(Generally, L is the easiest to compute first!)

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M/M/1

For the M/M/1 queueing system, then

$$W_q = W - \frac{1}{\mu} \Rightarrow W_q = \frac{\lambda}{\mu(\mu - \lambda)}$$

$$L_q = \lambda W_q \Rightarrow L_q = \frac{\lambda^2}{\mu(\mu - \lambda)}$$

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Example

An average of 24 trucks per 8-hour day arrive to be unloaded &/or loaded, which requires an average of 15 minutes.

The loading dock can handle only a single truck at a time.

Assume that the arrival process is Poisson, and that the service times have exponential distribution.

This loading dock is modeled as an M/M/1 queue.

M/M/1

λ = arrival rate = 3/hour

μ = service rate = 4/hour

Utilization of the server $\rho = \frac{\lambda}{\mu} = 0.75$

Average number of trucks in system $L = \frac{\rho}{1 - \rho} = \frac{0.75}{1 - 0.75} = 3$

Average time in system per truck $W = \frac{L}{\lambda} = \frac{3}{3/\text{hr}} = 1 \text{ hr.}$

Steady-state Behavior

M/M/1

λ = arrival rate = 3/hour

μ = service rate = 4/hour

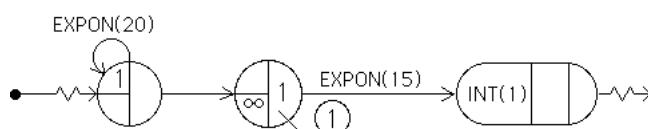
Average time in the queue $W_q = W - \frac{1}{\mu} = 1 \text{ hr.} - \frac{1}{4/\text{hr}} = 0.75 \text{ hr.}$

Average length of the queue $L_q = \lambda W_q = (3/\text{hr})(0.75\text{hr}) = 2.25$

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SLAM Model



M/M/1 Queue

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SLAM code

```

GEN,BRICKER,MM1_QUEUE,3/18/92,1,Y,Y,Y/N,Y,Y,72;
LIM,1,2,100;
INIT,,14400;           SIMULATE TEN 24-HOUR DAYS
NETWORK;
CREATE,EXPON(20),,1;   ARRIVAL OF TRUCKS
QUEUE(1);
ACTIVITY(1)/1,EXPON(15);
COLCT,INT(1),TIME IN SYSTEM,20/10/10;
TERM;
END;
FIN;

```

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STATISTICS FOR VARIABLES BASED ON OBSERVATION

FILE STATISTICS

TIME IN SYSTEM	MEAN VALUE	STANDARD DEVIATION	COEFF. OF VARIATION	MINIMUM VALUE	MAXIMUM VALUE	NO. OF OBS	FILE NUMBER	LABEL/ TYPE	AVERAGE LENGTH	STANDARD DEVIATION	MAXIMUM LENGTH	CURRENT LENGTH	AVERAGE WAIT TIME
	0.629E+02	0.650E+02	0.103E+01	0.430E-01	0.327E+03	741			2.499	3.640	20	1	48.431
							1	QUEUE	2.499	3.640	20	1	48.431
							2	CALENDAR	1.742	0.438	3	2	12.290

W
↑
↓

L_q

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SERVICE ACTIVITY STATISTICS

OBS FREQ	RELA FREQ	CELL LIM 0	20	40	60	80	100
122	0.165	0.100E+02	*****	+	+	+	+
103	0.139	0.200E+02	*****	C	C	C	C
70	0.094	0.300E+02	*****				
62	0.084	0.400E+02	****				
63	0.085	0.500E+02	****				
55	0.074	0.600E+02	****				
43	0.058	0.700E+02	****				
24	0.032	0.800E+02	***				
28	0.036	0.900E+02	***				
17	0.023	0.100E+03	**				
20	0.027	0.110E+03	**				
18	0.024	0.120E+03	**				
7	0.009	0.130E+03	*				
16	0.022	0.140E+03	*				
10	0.013	0.150E+03	*				
11	0.016	0.160E+03	*				
3	0.004	0.170E+03	*				
9	0.018	0.180E+03	*				
7	0.009	0.190E+03	*				
9	0.012	0.200E+03	*				
9	0.012	0.210E+03	*				
35	0.047	INF	***				
---			+	+	+	+	+
741			0	20	40	60	80
							100

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p
↑
↓

Characteristic		Observed in Simulation	Predicted by Theory
Utilization of server	ρ	74.2%	75%
Average time in system per truck	W	62.9 min.	60 min.
Average number in system	L_q	2.499	2.25
Average time in queue per truck	W_q	48.43 min.	45 min.

Example

The arrival process at a certain work center is Poisson, with rate 3/hour.

The service time has exponential distribution, with mean 0.3 hr.

Each job waiting for processing requires 1 m² of floor space.

How much in-process storage space should be allocated to accomodate all waiting jobs

... 90% of the time?

... 95% of the time?

... 99% of the time?

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If n square meters of floor space are allocated, this will be sufficient a fraction of time equal to

$$P\{\text{number in system} \leq n+1\} = \sum_{j=0}^{n+1} \pi_j$$

$$\pi_j = \rho^j (1 - \rho)$$

$$\begin{aligned} \Rightarrow \sum_{j=0}^{n+1} \pi_j &= \sum_{j=0}^{n+1} \rho^j (1 - \rho) \\ &= (1 - \rho) \sum_{j=0}^{n+1} \rho^j = (1 - \rho) \left[\frac{1 - \rho^{n+2}}{1 - \rho} \right] \\ &= 1 - \rho^{n+2} \end{aligned}$$

sum of finite geometric series

Solving for n : $1 - \rho^{n+2} \geq \alpha$

$$\Rightarrow \rho^{n+2} \leq 1 - \alpha$$

$$(n+2) \log \rho \leq \log(1 - \alpha)$$

$$n+2 \geq \frac{\log(1 - \alpha)}{\log \rho}$$

$$n \geq \frac{\log(1 - \alpha)}{\log \rho} - 2$$

& integer!

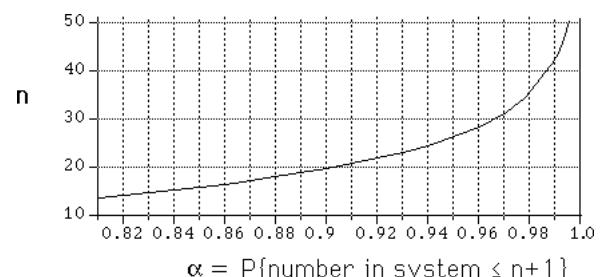
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$$n \geq \frac{\log(1-\alpha)}{\log \rho} - 2$$

where
 $\rho = 0.9$

α	$\frac{\log(1-\alpha)}{\log \rho} - 2$	n
0.85	16.00598614	17
0.9	19.85434533	20
0.95	26.43315881	27
0.96	28.5510637	29
0.97	31.28151799	32
0.98	35.12987717	36
0.99	41.70869065	42
0.992	43.82659554	44
0.994	46.55704984	47
0.996	50.40540902	51



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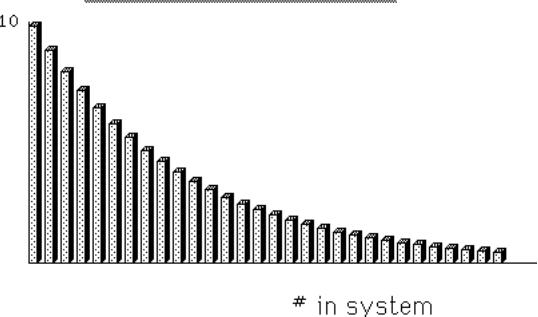
Steadystate Distribution

i	P _i	CDF
0	0.100000	0.100000
1	0.090000	0.190000
2	0.081000	0.271000
3	0.072900	0.343900
4	0.065610	0.409510
5	0.059049	0.468559
6	0.053144	0.521703
7	0.047830	0.569533
8	0.043047	0.612580
9	0.038742	0.651322
10	0.034868	0.686189
11	0.031381	0.717570
12	0.028243	0.745813
13	0.025419	0.771232
14	0.022877	0.794109
15	0.020589	0.814698

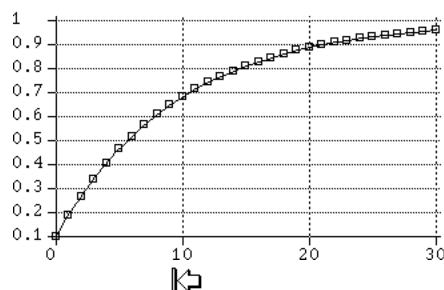
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i	P _i	CDF
15	0.020589	0.814698
16	0.018530	0.833228
17	0.016677	0.849905
18	0.015009	0.864915
19	0.013509	0.878423
20	0.012158	0.890581
21	0.010942	0.901523
22	0.009848	0.911371
23	0.008863	0.920234
24	0.007977	0.928210
25	0.007179	0.935380
26	0.006461	0.941850
27	0.005815	0.947665
28	0.005233	0.952899
29	0.004710	0.957609
30	0.004239	0.961848

Probability

Steadystate Distribution

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Cumulative Steady-State Probabilities

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